

# Behavioral Biases, Current Account and Default

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## Abstract

We include behavioral biases as in Barberis et al. [1998] in a real business cycle model with default and trend shocks. Agents infer the trend process using misspecified mental models. The misspecification responds to two known heuristics: representativeness and conservatism, that leads to behavioral responses from the agents in the form of waves of optimism, pessimism, overreaction and underreaction. We present suggestive evidence from IMF forecasts that these biases are indeed present. The model can explain stylized facts found in developing economies, namely, high consumption to output volatility, counter-cyclical trade balance and frequent defaults. The model is observationally equivalent to other proposed explanations for these effects, but with potentially different policy implications.

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# 1 Introduction

We include dynamic behavioral biases into an open economy real business cycle model with trend shocks and default as in Aguiar and Gopinath [2006]. We show that behavioral biases can account for features of developing economies such as debt crises, the excess volatility of consumption, and the counter-cyclical of the trade balance [Aguiar and Gopinath, 2006]. Agents endogenously move between waves of optimism and pessimism about the future trend. The model is observationally equivalent to Aguiar and Gopinath [2006] not due to external random trend shocks, but due to disciplined expectational mistakes.

We first present evidence from behavioral biases in IMF forecasts. Expanding on the methodology from Amir and Ganzach [1998], we show that IMF forecasts exhibit overoptimism and overpessimism, and that this depends on the sign of the forecast revision in previous periods. This is particularly striking for a group of middle income countries, including emerging Asia and Europe, and Latin America.

The expectation formation mechanism we adapt is from Barberis et al. [1998], which models investor sentiment parsimoniously and in a disciplined manner. In their model, the underlying asset follows a random walk, but the investor believes that the world can be in two regimes that alternate with some probability: mean-reverting and trend-following. Agents learn across these misspecified models to know which one is more likely to be behind the data generating process. Positive shocks followed by positive shocks increase the likelihood that the trend-following model is behind the process, and alternating shocks do the same for the mean reverting regime. As the agents learn through misspecified models, waves of optimism and pessimism arise, as well as overreaction and underreaction to news.

The underlying reason for agents to use misspecified models arises from the use of the use heuristics, specifically those of representativeness and conservatism. Representativeness, is part of a series of heuristics coined by Tversky and Kahneman [1974], in which people observe patterns in random sequences. This leads agents to quickly infer the underlying process on the basis of a small sample. Conservatism, identified by Edwards [1968], states that once agents form an impression, they are slow to change it when presented with new evidence. Agents using these heuristics, when they need to predict the

future realization of a random sequence of shocks, are prone to behave in waves of optimism and pessimism, overreaction and underreaction.

We extend this learning mechanism to a large state space and apply this to trend shocks in a dynamic stochastic general equilibrium model with default as in Eaton and Gersovitz [1981], Arellano [2008] and Aguiar and Gopinath [2006]. The model is able to explain the main features of developing economies' stylized facts. First, consumption is more volatile due to the agent expecting a higher wealth in the future that does not materialize. Second, the trade balance is countercyclical because periods of optimism lead to the agent to take debt on its permanent income. Finally, sudden stops occur when expectations suddenly change from one regime to another. This alternative explanation for the facts opens the possibility for macro-prudential policy to calm down expectations.

**Related Literature.** This paper deals with two branches of the literature: international macroeconomics in developing countries and behavioral biases and learning in macro.

The literature on equilibrium models for developing economies started with Mendoza [1991] and spanned a large literature. However, standard real business cycles models failed to account for features of the data in developing countries [Aguiar and Gopinath, 2006]. There are mainly two main types of explanations: those relying on non-stationary shocks [Aguiar and Gopinath, 2006] and those that rely on financial frictions [Garcia-Cicco et al., 2010, Chang and Fernández, 2013]. We provide an equilibrium model with stationary shocks and behavioral biases that can also account for those features.

Even though behavioral economics had a central role in financial literature, attempting to incorporate behavioral features in macroeconomics is relatively new [Akerlof, 2002, Hommes, 2021]. There is, however, a large literature on imperfect information and learning about the trend [Beaudry and Portier, 2006, Boz et al., 2011, Angeletos and La'O, 2013, Zeev et al., 2017, Ilut and Schneider, 2014, Cao and L'Huillier, 2018, Heymann et al., 2001]. In our paper “news” are, in a sense, endogenous and the agent learns in the space of (incorrect) mental models.

Closest to our model, Jaimovich and Rebelo [2007] present a macroeconomic model with overoptimism and overconfidence in a static manner. In contrast, we apply these insights to developing countries' issues and focus on the dy-

dynamic aspect of behavioral biases. Our expectation formation mechanism generates different beliefs depending on the state of the world.

There is also a branch of the literature testing for behavioral biases in macroeconomic data. Using state level data, Korniotis and Kumar [2011] find an effect of biases on the macroeconomy. Hagenhoff and Lustenhouwer [2020] identifies behavioral biases in professional forecasters. Ashiya [2003] studies overreaction and underreaction in Japanese forecasters. We adapt that methodology to a panel of predictions by the same forecaster, the IMF and find evidence of biases.

## 2 Some Evidence

There is ample statistical evidence on price overreaction and underreaction in financial markets, including exchange rates and commodity prices. [Park and Sohn, 2013, Barberis and Thaler, 2003, Frankel and Froot, 1990]. Using macroeconomic data Hagenhoff and Lustenhouwer [2020] and Ashiya [2003] find the presence of biases in professional forecasters. On the other hand, there is a large body of empirical and experimental evidence on the behavioral biases present in our paper [Baker and Nofsinger, 2010].

In this section, we provide suggestive evidence of the presence of behavioral biases in forecasts using IMF’s forecasts since 2008 until 2020. We adapt the methodology by Amir and Ganzach [1998] to study the presence of dynamic behavioral biases in IMF forecasts. The biases on which we focus are optimism, pessimism, overreaction and underreaction; behaviors that the expectation formation mechanism from Barberis et al. [1998] delivers as an outcome of agents using representativeness and conservatism heuristics.

**Methodology.** Identifying behavioral biases is challenging. Positive forecast errors can mean that the analyst is being optimistic or that it is overreacting to (positive) new information. Thus, a positive forecast error is compatible with both (i) an optimistic analyst and (ii) a pessimistic analyst overreacting to new information. A positive forecast revision and negative forecast error can also be equivalent to overreaction and pessimism. In the appendix, Figure 4 shows all possible observational outcomes of different biases.

In order to disentangle the effects, Amir and Ganzach [1998] regress forecast errors on forecast revisions. Let  $f_{t,i}^{t-2}$  be the forecast in  $t - 2$  for period  $t$  and

country  $i$ , and  $f_{t,i}^{t-1}$  the revised forecast in  $t - 1$  for  $t$  for the same country  $i$ . Let  $g_i^t$  be the realized growth rate of country  $i$ . Thus, the forecast error is  $FE_i^t \equiv f_{t,i}^{t-1} - g_i^t$  and the forecast revision is  $FR_i^t \equiv f_{t,i}^{t-1} - f_{t,i}^{t-2}$ .

$$FE_i^t = \alpha + \beta FR_i^t + \epsilon_i^t \quad (1)$$

When  $\alpha > 0$ , there is evidence of optimism, and when the reverse is true, there is evidence of pessimism. This is due to the fact that forecast errors should be uncorrelated with the past, and thus the constant term should be zero. On the other hand,  $\beta > 0$  implies overreaction to new information, and when the converse is true there is evidence of underreaction. In order to differentiate periods of good news and bad news, we split the sample depending on whether the forecast revisions in previous period  $j$  where positive ( $FR_{t-j}^+$ ) or negative ( $FR_{t-j}^-$ ).

Table 1 presents the estimates of (1) separated by the sign of the previous forecast revisions, for all countries and middle income countries (Emerging Asia, Latin America and Emerging Europe). When there were positive forecast revisions in the last two periods, the forecast is overoptimistic and underreacts to news for the case of middle income countries, but there is no such effect when countries are pooled together. For the case of bad news in the preceding two periods, there is no effect for any of the countries, but there is evidence of pessimism and underreaction for Middle Income countries. When focusing on only one period of positive forecast revisions, the same pattern is observed.

In the bottom part of Table 1 the results of the regression conditioning on alternating signs of forecast revisions in middle income countries is shown. All coefficients on the constant but one are statistically insignificant. In the case of a statistically positive coefficient (positive news followed by negative news in the previous periods) there is some evidence of optimism, which is consistent with the analyst expecting mean reversion. Notice that in this case there are several cases of overreaction. When news are alternating, there is overreaction to news, but when news are consistently positive or negative (Table 1), there is underreaction to good news.

These regressions show that there is some evidence suggestive of behavioral biases in IMF forecasts and particularly in middle income countries. Moreover, they exhibit dynamics effects depending on the sign of forecast revisions.

	$FR_{t-1}^+, FR_{t-2}^+$		$FR_{t-1}^-, FR_{t-2}^-$	
	All Countries	Middle Income	All countries	Middle Income
Constant	0.29 (0.38)	0.77* (0.46)	-1.44*** (0.29)	-1.3 (0.42)
FR	-0.20*** (0.05)	-0.14* (0.08)	-0.11*** (0.02)	-0.002 (0.03)

	$FR_{t-1}^+$		$FR_{t-1}^-$	
	All Countries	Middle Income	All countries	Middle Income
Constant	0.39 (0.31)	0.96** (0.46)	-1.33*** (0.35)	-0.92* (0.5)
FR	-0.05* (0.03)	0.066 (0.06)	-0.077*** (0.02)	0.09** (0.03)

Middle Income						
	$FR_{t-1}^+, FR_{t-2}^-$		$FR_{t-1}^-, FR_{t-2}^+$		$FR_{t-1}^-, FR_{t-2}^+, FR_{t-3}^-$	
Constant	-1.29 (1.44)		-1.17* (0.57)		-2.6 (2.4)	
FR	-0.38*** (0.09)		0.011* (0.06)		-0.34 (0.14)	

	$FR_{t-1}^-, FR_{t-2}^+, FR_{t-3}^+$		$FR_{t-1}^+, FR_{t-2}^-, FR_{t-3}^-$		$FR_{t-1}^+, FR_{t-2}^-, FR_{t-3}^-$	
Constant	-0.76 (0.53)		-1.01 (0.77)		-1.27 (1.7)	
FR	0.14* (0.06)		0.35** (0.12)		0.32** (0.14)	

Standard errors in parentheses. \*\*\*  $p < 0.01$ , \*\*  $p < 0.05$ , \*  $p < 0.1$

Table 1: Regression results. All regressions include year fixed effects.

### 3 The Model

#### 3.1 Expectations: two state example

This section presents the expectation formation mechanism in Barberis et al. [1998]. Suppose that the state of the world can be either  $y^+$  or  $y^-$ . The real process is *i.i.d.*, but the investor uses two mental models, which she believes to be the possible data generating processes. The first model,  $M1$ ,

$M1$	$y_{t+1} = y^+$	$y_{t+1} = y^-$
$y_{t+1} = y^+$	$\pi^H$	$1 - \pi^H$
$y_{t+1} = y^-$	$1 - \pi^H$	$\pi^H$
$M2$	$y_{t+1} = y^+$	$y_{t+1} = y^-$
$y_{t+1} = y^+$	$\pi^L$	$1 - \pi^L$
$y_{t+1} = y^-$	$1 - \pi^L$	$\pi^L$

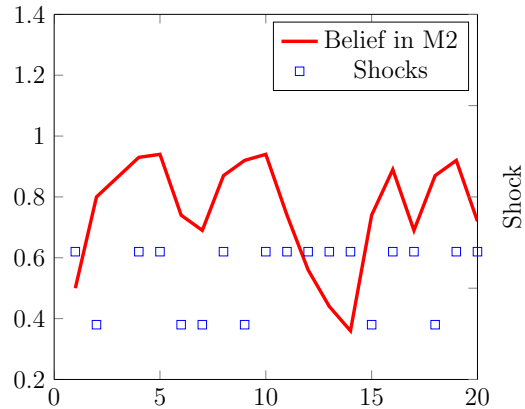


Figure 1: Left panel: Transition matrix for the two mental models,  $M1$  and  $M2$ . Right panel: beliefs (left axis) and shocks (right axis) across periods

is “trend-following” or high persistence. In this case, a positive shock is expected to be followed by a positive shock with a high probability  $\pi^H$ , and the same for negative shocks. The second model,  $M2$ , is “reversionist”, and expects positive shocks to be more likely succeeded by negative shocks, with probability  $1 - \pi^L$ . The transition matrices for the example are in Figure 1.

The investor believes that  $M1$  and  $M2$  switch as the data generating process with probability  $\gamma$ . In order to make decisions, the investor needs to forecast  $y$ . Thus, she attempts to assign probabilities  $\pi_t$  to each of the models as the underlying process. Thus, at any point  $t$ , she calculates  $\pi_t = Pr(M1^{t+1}|y_t, y_{t-1}, \pi_{t-1})$  using Bayes Rule. Barberis et al. [1998] show that this structure generates overreaction, underreaction, overconfidence and overoptimism.

In Figure 1 there is an example. The shock can be high or low and is *i.i.d.* However, the agent forms beliefs thinking one of the two models is the real one. In the first periods, shocks are alternating between high and low shocks, thus the probability of being in  $M2$ , the reversionist model, increases. Between periods 10 and 15, there are several periods of positive shocks, and thus the probability of being in the reversionist model plummets. The logic of the model is that agents tend to underreact, but after repeated observation of news, they may overreact to new information.

### 3.2 The Framework

The model uses the framework of Aguiar and Gopinath [2006], which itself expands Arellano [2008] with trend shocks. There is one country that has one period debt with the rest of the world, and if it defaults, it will be punished by an output cost and financial autarky for a period of time.

The economy has an endowment  $y_t$ , and the representative agent has constant risk aversion preferences given by

$$u = \frac{c^{1-\theta}}{1-\theta}$$

where  $\beta \leq 1$  is the discount factor. The endowment follows :

$$y_t = e^{z_0} \Gamma_t$$

where  $e^{z_0}$  is a constant, and  $\Gamma_t = g_t \Gamma_{t-1}$ , with the growth rate of the trend being represented by  $g_t$ . The true processes is given by

$$\ln g_t = (1 - \rho_g) (\ln(\mu_g) - c) + \rho_g \ln g_{t-1} + \epsilon_t^g$$

where  $|\rho_g| < 1$ ,  $\epsilon_t^g \sim N(0, \sigma_g^2)$  and  $c = \frac{1}{2} \frac{\sigma_g^2}{1-\rho_g^2}$ . There is no transitory shock. The process for the growth rate, implies that output will be permanent higher in all future periods. As in Aguiar and Gopinath [2006], we assume  $E \lim_{t \rightarrow \infty} \beta^t \Gamma_t^{1-\theta} = 0$  to ensure a well defined problem.

The agent has access to one period international bonds,  $a_t$ , that pay one unit of the consumption good. The price is  $q_t$ . At the beginning of the period the country decides whether to default or not. If it defaults, the value of such action is  $V^B$ . If the agent decides to repay the debt, the country will be in good credit standing and get  $V^G$ . Thus, the value of the agent is  $V(a_t, z_t, g_t) = \max\{V^G, V^B\}$ . If defaults, the agent will consume its endowment after paying a proportional penalty given by  $(1 - \delta)$ . The continuation value will depend on its redemption probability,  $\lambda$ .

The value of default is given by

$$V^B(0, g_t, g_{t-1}, \pi_{t-1}) = u((1-\delta)y_t) + \lambda \beta \tilde{E}_t V(0, g_{t+1}, g_t, \pi_t) + (1-\lambda) \tilde{E}_t V^B(0, g_{t+1}, g_t, \pi_t)$$

Where with probability  $\lambda$  the country is redeemed and continues with a clean slate. The exogenous states is  $g_t$ . Notice that given that the agent



uses previous information on the shocks to make inferences, the belief from previous period,  $\pi_{t-1}$  is also a state variable.  $\tilde{E}$  represents the fact that the agent is not using the true data generating process to make inferences.

The value of keeping a good credit standing is given by

$$V^G(a_t, g_t, g_{t-1}, \pi_{t-1}) = \max_{c_t} \left\{ u(c_t) + \beta \tilde{E}V(a_{t+1}, g_{t+1}, g_t, \pi_t) \right\}$$

$$s.t \quad c_t = y_t + a_t - q_t a_{t+1}$$

International investors are risk neutral and have an opportunity cost of  $r^*$ . They are aware of the agents problem and know that default,  $D(a_t, g_t) = 1$  will happen if  $V^B \geq V^G$ . Thus, they price the bond at

$$q_t(a_{t+1}, g_t) = \frac{\tilde{E}_t(1 - D_{t+1})}{1 + r^*}$$

Thus, the higher the expected probability of repayment, the higher the price of the bond. Notice that international investors share the same belief structure as domestic agents.

Agent believe there are two possible processes in governing the trend,

$$(M1) \quad \ln g_t = (1 - \rho_g^{M1})(\ln(\mu_g) - c) + \rho_g^{M1} \ln g_{t-1} + \epsilon_t^{M1}$$

$$(M2) \quad \ln g_t = (1 - \rho_g^{M2})(\ln(\mu_g) - c) + \rho_g^{M2} \ln g_{t-1} + \epsilon_t^{M2},$$

with persistence  $\rho_g^{M1} > \rho_g^{M2}$  and  $\epsilon_t^{Mi} \sim N(0, \sigma^{2, Mi})$  for  $i = 1, 2$ , where we allow the variance to be potentially misperceived. Finally, the agent believes the world switches across regimes with probability  $\gamma$ . Thus, the agent believes that the data generating process behind the observed sequence of shocks is  $M1$  with probability

$$\pi_{t+1} = \frac{(1 - \gamma)\pi_t + \gamma(1 - \pi_t)Pr(g_{t+1}|M1, g_t)}{((1 - \gamma)\pi_t + \gamma(1 - \pi_t))Pr(g_{t+1}|M1, g_t) + (\gamma\pi_t + (1 - \gamma)(1 - \pi_t))Pr(g_{t+1}|M2, g_t)}$$

A final point is regarding the issue of whether the agent understands that her forecast probabilities will change in the following period and thus the problem would not admit a simple recursive structure. Kreps [1998] shows a problem in which which an agent updates using Bayes' law but is myopic with respect to her future updating of probabilities. Kreps [1998] calls this "anticipated utility" and shows that is a good approximation to the rational expectations equilibrium.

### 3.3 Calibration and Model Solution

We calibrate the model exactly as Aguiar and Gopinath [2006] (AG, henceforth) in order to match Argentinian data. The values of the parameters are in Table 3 in Appendix B. The true data generating process has the same mean and variance as in AG, but without any autocorrelation,  $\rho^g = 0$ . We set  $M1$  to be the trend following model with  $\rho_g^{M1} = .11$  and  $\sigma_g^{2,M1} = 0.02$ , showing more precision in beliefs than in the true data generating process ( $\sigma_g^2 = 0.03$ ). In order to discipline the model, we choose values for  $M2$  such that on average, the processes  $M1$  and  $M2$  generate the data generating process. Finally, we set  $\gamma = .2$  so that regimes changes are rare. All lagged state variables,  $(s_{t-1}, \pi_{t-1})$  are summarized in the process of Bayesian updating. Our solution method discretizes the distribution,  $\pi \in [0, 1]$  and solves the agent's problem for each possible probability on the grid. Expectations are computed dynamically via Bayes' Rule when simulating the model and the relevant policy function is used in each state to obtain optimal responses.

## 4 Results

In Figure 2 we can observe a sample path of beliefs in  $M1$  (trend-following model with low variance) and shocks. In the left panel we can observe that in period 5 there is a negative shock but expectations do not react, and it is subsequently followed by higher shocks. However, in period 11 there is a large shock that induces a sudden reversal of expectations. Given the intermittent behavior of shocks afterwards, the probability oscillates between lower values and higher values. In the right panel, we can observe the slow build up of more optimistic expectations with a sequence of good shocks. During the simulations, the maximum value for the probability of  $M1$  is 0.92 and the minimum is 0.01.

In Figure 3, we can observe default regions under each mental model, if they were believed to be the true model with certainty. The slope for the default region under the trend following model,  $M1$  is more steep. In other words, the agent is less likely to default at higher levels of the shock even when debt is high, given that there is an expectation of the shock remaining higher. The excess default of the model with respect to a traditional model will be given by the expectational switch between the regions.

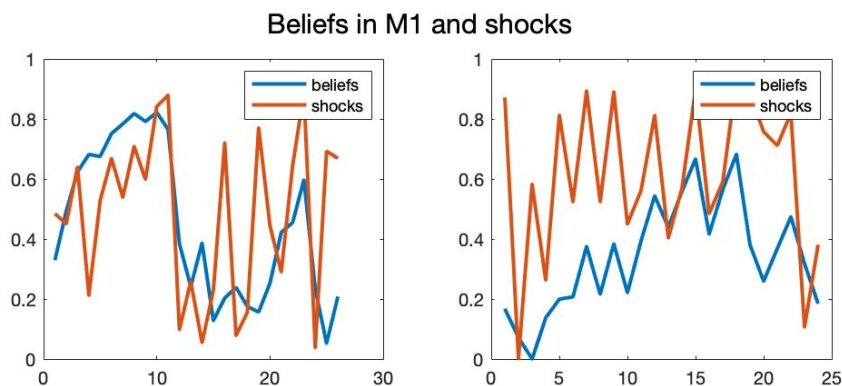


Figure 2: Sample of simulated beliefs (probability of  $M1$ ) and shocks. Shocks are normalized to be between 0 and 1.

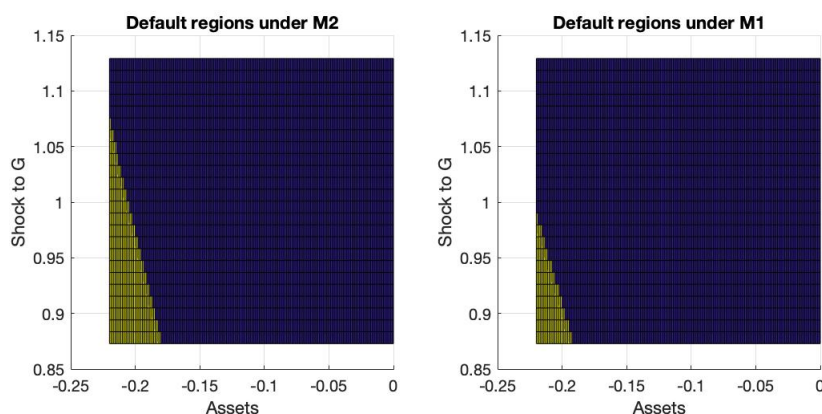


Figure 3: Default region for the case of pessimistic beliefs and optimistic beliefs. The yellow region represents combinations of the productivity state and assets for which the optimal choice is default. The blue shaded represents the non-default region. The vertical axis represents the realization of the shock. The horizontal axis represents assets normalized by (mean) trend income.

The results of the simulated economy are in Table 2. The expectationally augmented model replicates qualitatively the features of the data, including the puzzles in which AG focus on. This includes the excess volatility of consumption to output, a trade balance with larger variance than output, a countercyclical trade balance, and plausible frequency of defaults. The behavioral augmented model also explains to a better extent the dispersion

	Data	RBC Transitory	AG Trend Trend	Behavioral
$\sigma(y)$	4.08	4.32	4.45	4.14
$\sigma(c)$	4.85	4.37	4.71	4.46
$\sigma(TB/Y)$	1.36	0.17	0.95	1.17
$\sigma(R_s)$	3.17	0.04	0.32	2.25
$\rho(c, y)$	0.96	0.99	0.98	0.97
$\rho(TB/y, y)$	-0.89	-0.33	-0.19	-0.15
$\rho(R_s, y)$	-0.59	0.51	-0.03	0.01
$\rho(R_s, TB/Y)$	0.68	-0.21	0.11	0.30
Default per 10,000 quarters	75	2	23	53

Table 2: Results are averages over 500 simulations. Simulated data is HP-filtered as empirical data. Standard deviations are percentages.

of interest rates. The underlying dynamics is driven by the fact that the economy moves between waves of optimism into waves of pessimism even though the true data generating process has no persistence. Thus, after a sequence of good shocks, the agent believes the process to be persistently high. This persuades her to take debt in order to smooth out her future income. When expectations suddenly reverse, a debt crisis arises.

## 5 Discussion

There are two things that require discussion: the specificity of the model for developing economies and the possibility of learning the true process.

*Why developing countries?* There is evidence that poor people are more prone to behavioral biases than rich people (For example Bernheim et al. [2015]). Moreover, behavioral traits here depend on the intrinsic volatility of the process. Thus, innately more volatile countries would leave room for larger behavioral distortions. Finally, exchange rates and commodity prices have a documented behavior similar to financial instruments (for example Hsu et al. [2013]), and in middle income countries output is more related to swings in the exchange rate.

*Should learning rule this process out?* First, Barberis et al. [1998] perform

a simulation for the Bayes posterior odds ratio and they show that even after five years of data and and 80% of the time investors consider it just as good as the true model. Second, yet another behavioral bias can be present: confirmation bias [Rabin and Schrag, 1999]. This would limit the Bayesian updating in favor of slower learning. Finally, Weitzman [2007] rationalizes the equity premium puzzle with parameter uncertainty. In that context, Bayesian updating of unknown structural parameters adds a tail thickening effect to posteriors, and thus expectations are very sensitive to subjective prior beliefs even with asymptotically infinite data.

## 6 Conclusion

This paper has presented a parsimonious and disciplined model of behavioral biases embedded in a business cycle model with default. The model is able to replicate features of the data that are especially relevant for developing countries. The real data generating process is stationary (as Garcia-Cicco et al. [2010] find out with sufficiently long data), but the agents perceive it as going through different trends. This generates dynamics that are consistent with stylized facts. The effect relies on overconfidence as well (i.e, less perceived variance ) as consistent with the findings of Jaimovich and Rebelo [2007]. Policy implications are potentially different from a model with trend shocks: whereas in Aguiar and Gopinath [2006] the equilibrium is efficient, if the cause for cycles is due to behavioral misinterpretations, there is scope for macroprudential policy or countercyclical policies.

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## 7 Appendix

### A. Observational Equivalences

	Positive revision		Negative revision	
	Effect	Joint effect	Effect	Joint effect
Optimism & Over-reaction	+	+	+	?
	+		-	
Optimism & Under-reaction	+	?	+	+
	-		+	
Pessimism & Over-reaction	-	?	-	-
	+		-	
Pessimism & Under-reaction	-	-	-	?
	-		+	

Figure 4: Observational equivalence of some behavioral biases. Source Amir and Ganzach [1998].

## B. Calibration

General parameters		
Risk aversion	$\theta$	2
World interest rate	$r^*$	0.01
Output cost	$\delta$	0.02
Redemption prob.	$\lambda$	0.10
Mean growth rate	$\mu_g$	1.001
Transitory prod	-	NA
Trend prod		
Std Dev.	$\sigma_g$	0.032
Persistence.	$\rho_g$	0.0
Expectational parameters		
$\sigma_g^{M1}$		0.02
$\rho_g^{M1}$		0.11
$\sigma_g^{M2}$		0.03 Match DGP
$\rho_g^{M2}$		0.001 Match DGP
$\gamma$		0.3

Table 3: Calibration. Standard values are from AG. The true data generating process (DGP) does not display any autocorrelation. The parameters for model  $M2$  are set such that  $M1$  and  $M2$  behave as the DGP